



# Beijing-Dublin International College



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## SEMESTER I FINAL EXAMINATION - 2018/2019

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**School of INSERT SCHOOL:** Beijing University of Technology

**MODULE CODE and MODULE TITLE**

BDIC2005J/BDIC1033J, Probability and Statistics

**HEAD OF SCHOOL NAME:** BDIC

**MODULE COORDINATOR NAME\*:** Han Min, Zhao Xu, Min Hui

**OTHER EXAMINER NAME**

**Time Allowed: 90 minutes**

**Instructions for Candidates**

All questions carry equal marks. The distribution of marks in the right margin shown as a percentage gives an approximate indication of the relative importance of each part of the question.

**BJUT Student ID:** \_\_\_\_\_

**UCD Student ID:** \_\_\_\_\_

I have read and clearly understand the Examination Rules of both Beijing University of Technology and University College Dublin. I am aware of the Punishment for Violating the Rules of Beijing University of Technology and/or University College Dublin. I hereby promise to abide by the relevant rules and regulations by not giving or receiving any help during the exam. If caught violating the rules, I accept the punishment thereof.

**Honesty Pledge:** \_\_\_\_\_ **(Signature)**

**Instructions for Invigilators**

Non-programmable calculators are permitted.

No rough-work paper is to be provided for candidates.

Obtained score

**Question 1:****Vacancy (Each blank 3 marks)**

- (1) There are two events  $A$  and  $B$ .  $P(A)=0.1$ ,  $P(A \cup B)=0.4$ .

If  $A$  and  $B$  are mutually exclusive,  $P(B)=$  0.3.

If  $A$  and  $B$  are mutually independent,  $P(B)=$   $\frac{1}{3}$ .

- (2) Suppose the random variable  $X \sim U(a, b)$ .  $E(X)=5$ ,  $D(X)=3$ ,  $a=$  2,  $b=$  8.

- (3) Let  $X_1, X_2, \dots, X_n$  ( $n > 2$ ) be a sample from  $N(\mu, \sigma^2)$ .

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i, \quad S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2.$$

$$\sqrt{n}(\bar{X} - \mu) / \sqrt{S^2} \sim t_{n-1}, \quad (n-1)S^2 / \sigma^2 \sim \chi_{n-1}^2.$$

- (4) Let  $X_1, \dots, X_{10}$  be a sample from  $N(\mu, \sigma^2)$ . We can calculate the sample mean 5 and sample variance 0.25 are respectively.  $\mu$  and  $\sigma^2$  are unknown parameters.

The confidence interval of  $\mu$  with confidence level 95% is [4.82, 5.36].

The confidence interval of  $\sigma^2$  with confidence level 95% is [0.118, 2.833].

$$\begin{aligned} \mu: \frac{\bar{x} - \mu}{S/\sqrt{n}} &\sim t_{n-1} : \bar{x} \pm \frac{S}{\sqrt{n}} t_{n-1} \frac{\alpha}{2} = 5 \pm \frac{0.5}{\sqrt{10}} \cdot t_{9} \frac{0.05}{2} = 4.64 / 5.36 \\ \sigma^2: \frac{(n-1)S^2}{\sigma^2} &\sim \chi_{n-1}^2 : \frac{(n-1)S^2}{\chi_{n-1}^2(\frac{\alpha}{2})} = \frac{9 \times 0.25}{19.023} = 2.118 \end{aligned}$$

Obtained score

**Question 2: (14 marks)**

5 of the 8 guns have been calibrated, 3 have not been calibrated. The probability of a shooter hitting a target with a calibrated gun is 0.8; The probability of the shooter hitting target using an uncalibrated gun is 0.3.

- (a) Find the probability of the shooter not hitting target;  
(b) Now take one of the 8 guns and shoot, the result is that the shooter hits the target.

Find the probability that the gun used to shoot is a calibrated gun.

A: get calibrated gun      B: hit

$$P(\bar{B}) = P(\bar{B}|A) + P(\bar{B}|\bar{A}) = \frac{5}{8} \times 0.2 + \frac{3}{8} \times 0.7 = 0.3875$$

$$b) P(A|B) = \frac{P(A) - P(B|\bar{A})}{\frac{5}{8} \times 0.8 + \frac{3}{8} \times 0.3} = \frac{\frac{5}{8} \times 0.8}{\frac{5}{8} \times 0.8 + \frac{3}{8} \times 0.3} = 0.8163$$

Obtained score

**Question 3: (14 marks)**

Suppose that the probability density function of the continuous random variable  $X$  is

$$f(x) = \begin{cases} \frac{1}{c^2 + x^2}, & x \in (0, \frac{\pi}{4}); \\ 0, & \text{elsewhere.} \end{cases}$$

Find:

- (a) The value of constant  $c$ ;

- (b)  $P(\frac{\sqrt{3}\pi}{12} < X \leq \frac{\pi}{3})$ ;

- (c)  $E(X)$ .

$$a) F(1) = \int_0^{\frac{\pi}{4}} f(x) dx = 1 \Rightarrow \frac{1}{c} \arctan \frac{\pi}{c} = 1$$

$$c = \frac{\pi}{4}$$

$$f(x) = \frac{\pi}{4} \arctan \frac{4}{\pi} x$$

$$P(\frac{\sqrt{3}\pi}{12} < X \leq \frac{\pi}{3}) = F(\frac{\pi}{3}) - F(\frac{\sqrt{3}\pi}{12}) = \frac{\pi}{4} (\arctan \frac{4}{3} - \arctan \frac{\sqrt{3}}{3})$$

$$= 18 - 17$$

$$c) E(X) = \int_0^{\frac{\pi}{4}} x f(x) dx = \int_0^{\frac{\pi}{4}} \frac{x}{c^2 + x^2} dx = \frac{1}{2} \ln \frac{x^2 + c^2}{c^2} = \frac{1}{2} \ln \frac{\pi^2}{c^2}$$



Obtained  
score

#### Question 4: (14 marks)

Suppose that the joint probability density of two random variables  $X$  and  $Y$  are given by

$$f(x, y) = \begin{cases} c \cdot y^3(2-x), & 0 \leq x \leq 1, 0 \leq y \leq x, \\ 0, & \text{elsewhere.} \end{cases}$$

$\frac{3}{4/12}$

Find:

(a) The value of constant  $c$ ;

(b) The marginal densities  $f_X(x)$ ,  $f_Y(y)$ ;

(c) Whether the two random variables  $X$  and  $Y$  are independent or not?

(Please give your reason);

(d)  $E(XY)$

Handwritten solutions for Question 4:

$$1) F(x, y) = \int_0^x \int_0^y c y^3 (2-x) dy dx = 1 \Rightarrow c = \frac{3}{4}$$

$$f_X(x) = \int_0^x f(x, y) dy = \int_0^x \frac{3}{4} y^3 (2-x) dy = \frac{3}{4} x^4 (2-x)$$

$$f_Y(y) = \int_0^1 f(x, y) dx = \int_0^y \frac{3}{4} y^3 (2-x) dx = \frac{3}{4} y^3 (2-y)$$

$$E(XY) = \int_0^1 \int_0^x \frac{3}{4} x y^4 (2-x) dy dx = \frac{21}{40}$$

Conclusion:  $f_X(x) \neq f_X(x) \cdot f_Y(y)$  → Not independent

Obtained  
score

#### Question 5: (14 marks)

Suppose  $X_1, X_2, \dots, X_n$  is a sample from  $X$ , and the probability density function of  $X$  is

$$f(x, \theta) = \begin{cases} \theta^2 x e^{-\theta x}, & x > 0, \\ 0, & x \leq 0, \end{cases}$$

where  $\theta > 0$ . Find:

(a) The moment estimator of  $\theta$ ;

(b) The maximum likelihood estimator of  $\theta$ .

Handwritten solutions for Question 5:

$$1) E(X) = \int_0^\infty x f(x) dx = \int_0^\infty \theta^2 x^2 e^{-\theta x} dx = \frac{2}{\theta}$$

$$2) L(\theta) = \prod_{i=1}^n f(x_i) = \theta^{2n} \cdot x_i! \cdot e^{-\theta \sum_{i=1}^n x_i}$$

$$\ln L(\theta) = 2n \ln \theta + \ln x_i! - \theta \sum_{i=1}^n x_i$$

$$\frac{\partial \ln L(\theta)}{\partial \theta} = \frac{2n}{\theta} - \sum_{i=1}^n x_i = 0 \Rightarrow \hat{\theta} = \frac{2}{\bar{x}}$$

Obtained  
score

#### Question 6: (14 marks)

The capacities (in ampere-hours) of 10 batteries were recorded as follows:

102, 97, 101, 103, 101, 98, 99, 104, 103, 98

Under the assumption that the capacity is normal distributed as  $N(\mu, \sigma^2)$ .

Question: at level of significance 0.05,

(a) test  $H_0: \mu = 100 \leftrightarrow H_1: \mu \neq 100$ ,

(b) test  $H_0: \sigma^2 = 2.5 \leftrightarrow H_1: \sigma^2 > 2.5$ .

The  $t$  distribution table and the  $\chi^2$  distribution table

$t_9(0.025) = 2.2622$	$t_9(0.05) = 1.8331$	$t_{10}(0.025) = 2.2281$	$t_{10}(0.05) = 1.8125$
$\chi_9^2(0.025) = 19.023$	$\chi_9^2(0.05) = 16.919$	$\chi_9^2(0.975) = 2.700$	$\chi_9^2(0.95) = 3.325$
$\chi_{10}^2(0.025) = 20.483$	$\chi_{10}^2(0.05) = 18.307$	$\chi_{10}^2(0.975) = 3.247$	$\chi_{10}^2(0.95) = 3.940$

Handwritten calculations for Question 6(a):

$$\bar{x} = 100.6$$

$$s^2 = 5.34$$

$$\frac{|\bar{x} - \mu|}{s/\sqrt{n}} = 0.82$$

$$t_{p, 0.025} = 2.2622 > 0.82 \text{ acceptable}$$

Handwritten calculations for Question 6(b):

$$\frac{(n-1)s^2}{\sigma_0^2} = \frac{9 \times 5.34}{2.5} = 19.224$$

$$\chi_{n-1}(2) = \chi_{9, 0.05} = 16.919$$

$$\chi_{n-1}(2) < 19.224$$

SO reject

**Appendix :**